

TWO PARAMETRIC GENERALIZATIONS OF HARVDA-CHARVAT'S ENTROPY AND ITS APPLICATION IN SOURCE CODING

TABASUM FATIMA¹, MIRZA ABDUL KHALIQUE BAIG¹

Manuscript received: 03.07.2025; Accepted paper: 24.11.2025;

Published online: 30.12.2025.

Abstract. In this manuscript, we present a new two-parametric generalization of Harvda-Charvat's measure of entropy $H_\alpha^\beta(P)$ and its salient characteristics. We also obtain its most significant entropies that are widely known and have a sway in the literature of information and coding theory. Furthermore, we also present a new generalized mean code-word length $L_\alpha^\beta(P)$, and we additionally ascertain how $H_\alpha^\beta(P)$ and $L_\alpha^\beta(P)$ are interrelated in terms of the source coding theorem and also demonstrate it through the Shannon-Fano Coding Algorithm. Finally, we check its monotonicity through a dataset on precipitation.

Keywords: Shannon's entropy; Harvda-Charvat's entropy; Average length; L'Hospital's rule.

1. INTRODUCTION

Information Theory is a mathematical field founded by Claude E. Shannon [1] aimed at analyzing the mechanisms involved in the transmission, storage, and measurement of information. It establishes systematic techniques for assessing how effectively messages can be communicated from a source to a destination through a communication channel, while accounting for challenges such as noise, interference, and signal degradation. Consider a discrete random variable $X = \{x_1, x_2, x_3, \dots, x_n\}$ with its respective probabilities $P = \{p_1, p_2, p_3, \dots, p_n\}$, then the concept of entropy is defined as:

$$H(P) = - \sum_{i=1}^n p_i \log_D p_i \quad (1)$$

The unit of entropy is taken to the base of the logarithm D , if $D = 2$, then entropy measure is known as a bit; $D = e$, then entropy measure is known as Nat; and if $D = 10$, then entropy measure is known as Hartley. Numerous generalized measures of Shannon's entropy under a discrete random variable have been presented in the literature of information theory. Harvda-Charvat [2] presented an idea of parametric entropy and defined the entropy of order β as:

$$H^\beta(P) = \frac{1}{1-\beta} [\sum_{i=1}^n p_i^\beta - 1], \beta > 0, \beta \neq 1 \quad (2)$$

¹ University of Kashmir, Department of Statistics, 190006 Srinagar, India. E-mail: tabasumf11@gmail.com; baigmak@gmail.com.

Apart from Harvda-Charvat [2], various other researchers, viz., Rényi [3], Campbell, L. L. [4], Sharma and Mittal [5], Hooda, D. S., Bhaker, U. S. [6], Bhat and Baig [7-12], etc., have also developed some generalized measures in the theory of information.

2. A NEW TWO-PARAMETRIC GENERALIZATION OF HARVDA-CHARVAT'S MEASURE OF ENTROPY

Let $X = \{x_1, x_2, x_3, \dots, x_n\}$ with their respective probabilities $P = \{p_1, p_2, p_3, \dots, p_n\}$ then we derived a new two-parametric generalization of Harvda-Charvat's entropy $H_\alpha^\beta(P)$ is given by

$$H_\alpha^\beta(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right], \alpha > 0, \beta > 0, \alpha \neq \beta \quad (3)$$

where α and β represent scaling factors that adjust how the entropy responds to different probability distributions. Let us now explicate α and β . From the application point

- Harvda-Charvat's entropy can be tuned to reduce the influence of rare or extreme events, unlike Shannon's entropy, which is very sensitive to small probabilities.
- Moreover, it is used to enhance signal extraction in a noisy environment by reducing the effect of random outliers.

2.1. PARTICULAR CASES OF OUR PROPOSED GENERALIZED HARVDA-CHARVAT'S MEASURE OF ENTROPY

a. When $\alpha = 1$, equation (3) reduces to Harvda-Charvat's [2] entropy of order β i.e.,

$$H_{\alpha=1}^\beta(P) = H^\beta(P) = \frac{1}{1-\beta} \left[\sum_{i=1}^n p_i^\beta - 1 \right]$$

b. When $\beta = 1$, equation (3) reduces to Harvda-Charvat's [2] entropy of order $\frac{1}{\alpha}$ i.e.,

$$H_\alpha^{\beta=1}(P) = H_{\frac{1}{\alpha}}(P) = \frac{1}{1-\frac{1}{\alpha}} \left[\sum_{i=1}^n p_i^{\frac{1}{\alpha}} - 1 \right]$$

c. When $\beta = 2\alpha$, equation (3) reduces to Harvda-Charvat's quadratic entropy, i.e.

$$H^{\beta=2\alpha}(P) = H^2(P) = 1 - \sum_{i=1}^n p_i^2$$

d. When $\alpha = 1$ and $\beta \rightarrow 1$, then by applying L'Hospital's rule, equation (3) reduces to the entropy given by Shannon [1], i.e.,

$$H_{\alpha=1}^{\beta \rightarrow 1}(P) = H(P) = - \sum_{i=1}^n p_i \log_D p_i$$

e. When $\beta = 1$ and $\alpha \rightarrow 1$, then applying L'Hospital's rule, equation (3) reduces to the entropy given by Shannon [1], i.e.,

$$H_{\alpha \rightarrow 1}^{\beta=1}(P) = H(P) = - \sum_{i=1}^n p_i \log_D p_i$$

f. When $\beta \rightarrow \alpha$, then by applying L'Hospital's rule, equation (3) reduces to the entropy given by Shannon [1], i.e.,

$$H^{\beta \rightarrow \alpha}(P) = H(P) = - \sum_{i=1}^n p_i \log_D p_i$$

g. When $\beta > 0$, $\alpha > 0$ and $\beta \neq \alpha$, and if all the events are equally likely, i.e., $p_i = \frac{1}{n}, \forall i = 1, 2, 3, \dots, n$, then we have

$$H_{\alpha}^{\beta}\left(\frac{1}{n}\right) = H\left(\frac{1}{n}\right) = \log_D n,$$

which is maximum entropy.

2.2. PROPERTIES OF OUR PROPOSED GENERALIZED HARVDA-CHARVAT'S MEASURE OF ENTROPY

Some significant features of our generalized entropy measure $H_{\alpha}^{\beta}(P)$ have been scrutinized in this section:

Property 1. $H_{\alpha}^{\beta}(P) > 0$, when $\alpha > 0$ and $\beta > 0$.

Proof: We have

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right], \alpha > 0, \beta > 0, \neq \alpha.$$

Case i: For $\beta > \alpha$.

For $\beta > \alpha$, we have $\frac{\beta}{\alpha} > 1$. Since $0 \leq p_i \leq 1, \forall i = 1, 2, 3, \dots, n$ and $\sum_{i=1}^n p_i = 1$, which implies that

$$p_i^{\frac{\beta}{\alpha}} < p_i$$

After some mathematical operation, it follows that:

$$\left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] < 0 \quad (4)$$

We have $\beta > \alpha$, which implies $\beta - \alpha < 0$. Also, for $\alpha > 0$, so we have

$$\frac{\alpha}{\alpha - \beta} < 0 \quad (5)$$

On combining equations (4) and (5), we have for $\beta > \alpha$

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] > 0 \quad (6)$$

Case ii: When $\beta < \alpha$.

When $\beta < \alpha$, we have $\frac{\beta}{\alpha} < 1$. Since, $0 \leq p_i \leq 1, \forall i = 1, 2, 3, \dots, n$ and $\sum_{i=1}^n p_i = 1$, which implies that

$$p_i^{\frac{\beta}{\alpha}} > p_i$$

After some mathematical operation, it follows that:

$$\left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] > 0 \quad (7)$$

As we have $\beta < \alpha$, which implies that $\alpha - \beta > 0$. Also for $\beta > 0$, so we have

$$\frac{\alpha}{\alpha - \beta} > 0 \quad (8)$$

On combining equation (7) and (8), we get

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] > 0 \quad (9)$$

From equations (6) and (9), we noticed that $H_{\alpha}^{\beta}(P)$ is positive for the defined values of the parameters α and β i.e.,

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] > 0$$

For $\beta > 0, \alpha > 0, \beta \neq \alpha$.

Property 2. $H_{\alpha}^{\beta}(P)$ is a symmetric function on every $p_i, i = 1, 2, 3, \dots, n$

Proof: This property is obviously true, i.e.,

$$H_{\alpha}^{\beta}(p_1, p_2, \dots, p_{n-1}, p_n) = H_{\alpha}^{\beta}(p_n, p_1, p_2, \dots, p_{n-1})$$

Property 3. The maximum value of $H_{\alpha}^{\beta}(P)$ is achieved when the choice of the occurrence of all the events is equal.

Proof: We have

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right], \beta > 0, \alpha > 0, \beta \neq \alpha$$

Assume the choice of the occurrence of all the events is equal, i.e., $p_i = \frac{1}{n}, \forall i = 1, 2, 3, \dots, n$, then we have

$$H_\alpha^\beta(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n \left(\frac{1}{n} \right)^{\frac{\beta}{\alpha}} - 1 \right]$$

After some mathematical computations, it follows that

$$H_\alpha^\beta \left(\frac{1}{n} \right) = H \left(\frac{1}{n} \right) = \log_D n$$

Which is the maximum entropy.

Property 4. The additive property is satisfied by $H_\alpha^\beta(P)$ in the following mathematical perspective:

$$H_\alpha^\beta(P * Q) = H_\alpha^\beta(P) + H_\alpha^\beta(Q)$$

where,

$$(P * Q) = \{p_1 q_1, \dots, p_1 q_m, p_2 q_1, \dots, p_n q_1, \dots, p_n q_m\}$$

is the joint probability mass function of two independent discrete random variables.

Proof: Suppose $(P * Q) = \{p_1 q_1, \dots, p_1 q_m, p_2 q_1, \dots, p_n q_1, \dots, p_n q_m\}$, be the joint probability mass function of two independent discrete random variables, then

We have

$$H_\alpha^\beta(P * Q) = \frac{\alpha}{\alpha - \beta} \left[\left(\sum_{i=1}^n \sum_{j=1}^m (p_i q_j)^{\frac{\beta}{\alpha}} - 1 \right) \right] \quad (10)$$

We can rewrite $(p_i q_j)^{\frac{\beta}{\alpha}}$ as

$$(p_i q_j)^{\frac{\beta}{\alpha}} = p_i^{\frac{\beta}{\alpha}} q_j^{\frac{\beta}{\alpha}}$$

Thus, the summation becomes

$$\sum_{i=1}^n \sum_{j=1}^m (p_i q_j)^{\frac{\beta}{\alpha}} = \sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \sum_{j=1}^m q_j^{\frac{\beta}{\alpha}}$$

The entropy

$$H_{\alpha}^{\beta}(P) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right]$$

and

$$H_{\alpha}^{\beta}(Q) = \frac{\alpha}{\alpha - \beta} \left[\sum_{j=1}^m q_j^{\frac{\beta}{\alpha}} - 1 \right]$$

On simplifying the above mathematical computation, we have

$$H_{\alpha}^{\beta}(P * Q) = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] + \frac{\alpha}{\alpha - \beta} \left[\sum_{j=1}^m q_j^{\frac{\beta}{\alpha}} - 1 \right]$$

Which implies,

$$H_{\alpha}^{\beta}(P * Q) = H_{\alpha}^{\beta}(P) + H_{\alpha}^{\beta}(Q)$$

This completes the proof.

3. SOURCE CODING THEOREMS

Consider a finite input source symbol $X = \{x_1, x_2, x_3, \dots, x_n\}$ with their respective probabilities of transmission $P = \{p_1, p_2, p_3, \dots, p_n\}$. Suppose we have code-words which have lengths $l_1, l_2, l_3, \dots, l_n$ and then the expected length of the coded message is defined by Shannon [1] as:

$$L(P) = \sum_{i=1}^n p_i l_i \quad (11)$$

A code is said to be a uniquely decipherable code over an alphabet of D symbols with length $L = \{l_1, l_2, l_3, \dots, l_n\}$ if and only if the Kraft's inequality holds, i.e.,

$$\sum_{i=1}^n D^{-l_i} \leq 1 \quad (12)$$

For all codes satisfying the inequality (12), the mean code-word length $L(P)$ defined at (11), lies between $H(P)$ and $H(P) + 1$ i.e.,

$$H(P) < L(P) < H(P) + 1$$

This is also called Shannon's noiseless coding theorem. Kapur [13] defined his mean code-word length for a discrete channel as:

$$L^\beta(P) = \frac{1}{1-\beta} \left\{ \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-1}{\beta} \right)} \right]^\beta - 1 \right\}, \beta > 0, \beta \neq 1 \quad (13)$$

and also showed that $L^\beta(P)$ lies between $H^\beta(P)$ and $H^\beta(P) + 1$ under the condition if the codes satisfy inequality (12), i.e.,

$$H^\beta(P) < L^\beta(P) < H^\beta(P) + 1$$

Numerous generalized source coding theorems under the condition of unique decipherability have been introduced by numerous scholars over the past few decades, for example, publications [14-16].

We presented a new generalized mean code-word length $L_\alpha^\beta(P)$ in this manuscript as:

$$L_\alpha^\beta(P) = \frac{\alpha}{\alpha-\beta} \left\{ \left[\left(\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-\alpha}{\beta} \right)} \right) \right]^\beta - 1 \right\}, \alpha > 0, \beta > 0, \alpha \neq \beta \quad (14)$$

where D is the number of alphabets used to code input source symbols.

3.1. PARTICULAR CASES OF NEW GENERALIZED MEAN CODE-WORD LENGTH

Case i: When $\alpha = 1$, (14) reduces to code-word length corresponding to Kapur's mean code-word length, i.e.,

$$L_{\alpha=1}^\beta(P) = L^\beta(P) = \frac{1}{1-\beta} \left\{ \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-1}{\beta} \right)} \right]^\beta - 1 \right\} \quad (15)$$

Case ii: When $\alpha = 1$, and $\beta \rightarrow 1$, then, by applying L'Hospital's rule (14) reduces to the optimal mean code-word length corresponding to Shannon entropy i.e.,

$$L_{\alpha=1}^{\beta \rightarrow 1}(P) = L(P) = \sum_{i=1}^n p_i l_i$$

Case iii: When $\alpha \rightarrow \beta$, then, by applying L'Hospital's rule (14) reduces to the optimum mean code-word length given by Shannon [1], i.e.,

$$L^{\alpha \rightarrow \beta} = L(P) = \sum_{i=1}^n p_i l_i$$

Now we will derive the relationship between (3) and (14) in terms of source coding theorems.

3.2. RELATIONSHIP OF HARVDA-CHARVAT'S GENERALIZED MEASURE AND ITS CORRESPONDING CODE-WORD LENGTH

The two theorems below show the relationship between the generalized measure and its corresponding code-word length.

Theorem 1. For all alphabets with $D > 1$ symbols, suppose the set of code-word lengths is $L = \{l_1, l_2, l_3, \dots, l_n\}$, which satisfy Kraft's inequality. The relationship between $H_\alpha^\beta(P)$ and $L_\alpha^\beta(P)$ is given by:

$$H_\alpha^\beta(P) \leq L_\alpha^\beta(P)$$

and equality, i.e., $H_\alpha^\beta(P) = L_\alpha^\beta(P)$ holds if and only if

$$l_i = -\log_D \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right] \quad (16)$$

Proof: For all $x_i, y_i > 0, i = 1, 2, 3, \dots, n$ and $\frac{1}{\theta} + \frac{1}{\delta} = 1, \theta < 1 (\neq 0), \delta < 1 (\neq 0), \theta < 0$, then by the reverse of Holder's inequality, we have

$$\left(\sum_{i=1}^n x_i^\theta \right)^{\frac{1}{\theta}} \left(\sum_{i=1}^n y_i^\delta \right)^{\frac{1}{\delta}} \leq \left(\sum_{i=1}^n x_i y_i \right) \quad (17)$$

The equality of (17) holds if $\exists c > 0$, such that

$$x_i^\theta = c y_i^\delta \quad (18)$$

Let

$$x_i = p_i^{\frac{\beta}{\beta-\alpha}} D^{-l_i} \quad y_i = p_i^{\frac{\beta}{\alpha(\beta-\alpha)}}$$

$$\theta = \frac{\beta - \alpha}{\beta} \quad \delta = \beta - \alpha$$

Substituting the above values in (17) and after some mathematical calculation, we get

$$\left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-\alpha}{\beta} \right)} \right]^{\frac{\beta}{\beta-\alpha}} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right]^{\frac{1}{\beta-\alpha}} \leq \sum_{i=1}^n D^{-l_i}$$

By using inequality (17) and after some mathematical calculation, we get

$$\left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right]^{\frac{1}{\beta-\alpha}} \leq \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-\alpha}{\beta} \right)} \right]^{\frac{\beta}{\beta-\alpha}} \quad (19)$$

Again, after some mathematical calculations on inequality (19) and subtracting 1 from both sides, we get

$$\frac{\alpha}{\beta - \alpha} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right] \leq \frac{\alpha}{\beta - \alpha} \left\{ \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta - \alpha}{\beta} \right)} \right]^{\beta} - 1 \right\}$$

Or we can rewrite the above inequality as:

$$H_{\alpha}^{\beta}(P) \leq L_{\alpha}^{\beta}(P)$$

Now we will show the equality i.e., $H_{\alpha}^{\beta}(P) = L_{\alpha}^{\beta}(P)$ holds if and only if

$$l_i = -\log_D \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right]$$

After some mathematical calculations, we get

$$D^{-l_i} = \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right]$$

Now, after multiplying throughout by p_i to the above equation, and applying appropriate mathematical calculations, it follows that:

$$\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta - \alpha}{\beta} \right)} = \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right]^{\frac{1}{\beta}} \quad (20)$$

After some mathematical calculations, we have

$$\frac{\alpha}{\beta - \alpha} \left\{ \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta - \alpha}{\beta} \right)} \right]^{\beta} - 1 \right\} = \frac{\alpha}{\alpha - \beta} \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} - 1 \right]$$

Or we can rewrite the above equality as

$$H_{\alpha}^{\beta}(P) = L_{\alpha}^{\beta}(P)$$

Theorem 2. For a code-word with lengths $L = \{l_1, l_2, l_3, \dots, l_n\}$ satisfying Kraft's inequality, and then $H_{\alpha}^{\beta}(P)$ and $L_{\alpha}^{\beta}(P)$ are related as follows:

$$L_{\alpha}^{\beta}(P) < H_{\alpha}^{\beta}(P) D^{\alpha - \beta} + \frac{\alpha}{\alpha - \beta} (D^{\alpha - \beta} - 1), \text{ when } \alpha > 0, \beta > 0, \beta \neq \alpha.$$

Proof: From Theorem 1, we see that $H_{\alpha}^{\beta}(P) = L_{\alpha}^{\beta}(P)$ is satisfied if and only if

$$l_i = -\log_D \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right]$$

The above expression can also be written as

$$l_i = -\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right]$$

Let code-word lengths $L = \{l_1, l_2, l_3, \dots, l_n\}$ be such that they satisfy the following inequalities:

$$-\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] \leq l_i < -\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] + 1 \quad (21)$$

Consider the interval

$$\theta_i = \left[-\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right], -\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] + 1 \right]$$

of length 1. In every θ_i , there lies exactly one positive integral l_i , such that

$$0 < -\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] \leq l_i < -\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] + 1 \quad (22)$$

We will first show that sequence $l_1, l_2, l_3, \dots, l_n$, defined satisfies the Kraft's inequality. From the left-hand side of (23), we have

$$-\log_D p_i^{\frac{\beta}{\alpha}} + \log_D \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] \leq l_i$$

or equivalently

$$D^{-l_i} \leq \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right] \quad (23)$$

Taking the summation over $i = 1, 2, 3, \dots, n$ on both sides of (23), we have

$$\sum_{i=1}^n D^{-l_i} \leq 1$$

This is Kraft's (1949) inequality. Now the last inequality of (23) gives:

$$D^{l_i} < \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right]^{-1} D \quad (24)$$

After some mathematical calculations, we have

$$D^{l_i \left(\frac{\alpha-\beta}{\beta} \right)} < \left[\frac{p_i^{\frac{\beta}{\alpha}}}{\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}}} \right]^{\frac{\beta-\alpha}{\beta}} D^{\frac{\alpha-\beta}{\beta}} \quad (25)$$

Now multiplying inequality (25) both sides by p_i , then summing over $i = 1, 2, 3, \dots, n$ and after suitable simplifications, we have

$$\sum_{i=1}^n p_i D^{l_i \left(\frac{\alpha-\beta}{\beta} \right)} < \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right]^{\frac{1}{\beta}} D^{\frac{\alpha-\beta}{\beta}} \quad (26)$$

Raising both sides to the power β and subtracting 1 throughout the inequality (26), we have

$$\frac{\alpha}{\alpha-\beta} \left\{ \left[\sum_{i=1}^n p_i D^{-l_i \left(\frac{\beta-\alpha}{\beta} \right)} \right]^{\beta} - 1 \right\} < \frac{\alpha}{\alpha-\beta} \left\{ \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] D^{\alpha-\beta} - 1 \right\} \quad (27)$$

From right hand side of (27) we have

$$\begin{aligned} & \frac{\alpha}{\alpha-\beta} \left\{ \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] D^{\alpha-\beta} - 1 \right\} \\ &= \frac{\alpha}{\alpha-\beta} \left\{ \left[\sum_{i=1}^n p_i^{\frac{\beta}{\alpha}} \right] - 1 \right\} D^{\alpha-\beta} + \frac{\alpha}{\alpha-\beta} (D^{\alpha-\beta} - 1) \end{aligned}$$

or we can write that

$$L_{\alpha}^{\beta}(P) < H_{\alpha}^{\beta}(P) D^{\alpha-\beta} + \frac{\alpha}{\alpha-\beta} (D^{\alpha-\beta} - 1).$$

Hence from above two coding theorems, we conclude that

$$H_{\alpha}^{\beta}(P) \leq L_{\alpha}^{\beta}(P) < H_{\alpha}^{\beta}(P) D^{\alpha-\beta} + \frac{\alpha}{\alpha-\beta} (D^{\alpha-\beta} - 1)$$

when $\alpha > 0, \beta > 0, \alpha \neq \beta$.

4. SHANNON-FANO CODING ALGORITHM

In this portion, we will demonstrate the validity of Theorems 1 and 2 by taking empirical data from [17] given in Table 1. The probability values of p_1, p_2, p_4, p_5 are given as 0.25, 0.5, 0.125, 0.125 respectively. Now, by using the Shannon-Fano Coding Algorithm, we have got the binary codes and code-word lengths corresponding to each probability given in Table 1.

Table 1. Shannon-Fano Coding Algorithm.

| p_i | Shannon-Fano Codes | l_i | α | β | $H_\alpha^\beta(P)$ | $L_\alpha^\beta(P)$ | $D^{\alpha-\beta}$ | $(D^{\alpha-\beta} - 1)$ | $H_\alpha^\beta(P)D^{\alpha-\beta} + \frac{\alpha}{\alpha-\beta}(D^{\alpha-\beta} - 1)$ |
|-------|--------------------|-------|----------|---------|---------------------|---------------------|--------------------|--------------------------|---|
| 0.25 | 001 | 3 | 0.9 | 0.6 | 1.5603 | 2.4527 | 1.2311 | 0.6933 | 2.6142 |
| 0.5 | 10 | 2 | | | | | | | |
| 0.125 | 1101 | 4 | | | | | | | |
| 0.125 | 1111 | 4 | | | | | | | |

Now the table allows us to deduce that the validity of Theorem 1 and Theorem 2 extends to the Shannon-Fano coding scheme as $H_\alpha^\beta(P) \leq L_\alpha^\beta(P)$ and $H_\alpha^\beta(P) \leq L_\alpha^\beta(P) < H_\alpha^\beta(P)D^{\alpha-\beta} + \frac{\alpha}{\alpha-\beta}(D^{\alpha-\beta} - 1)$ when, $\alpha > 0, \beta > 0, \alpha \neq \beta$.

5. REAL LIFE APPLICATION

Hinkley [18] provided a dataset comprising thirty consecutive years of March precipitation measurements, recorded in inches, which has also been used by [19]. The dataset captures variations in rainfall over this period and serves as a valuable resource for statistical and climatological analysis. The recorded values are as follows:

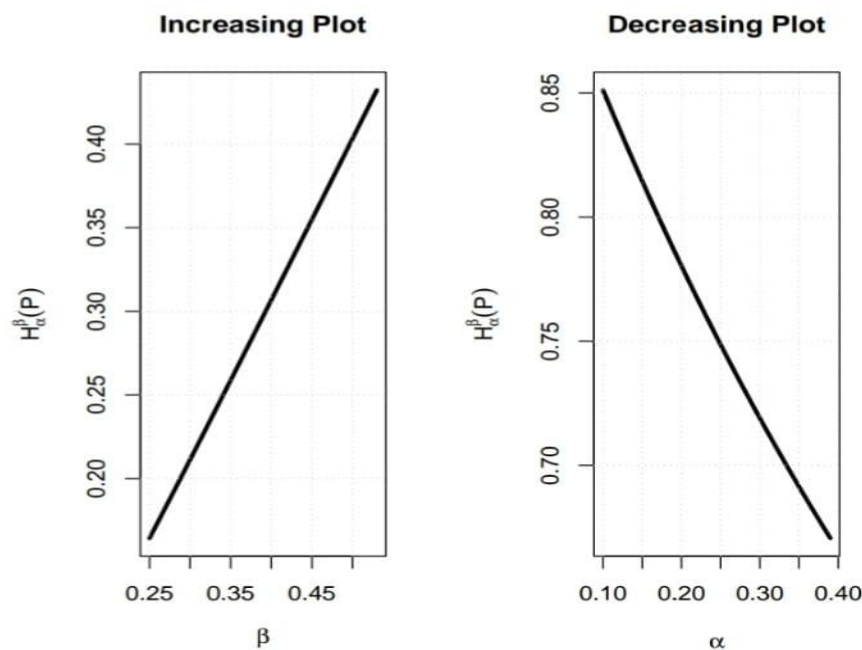
0.77 1.74 0.81 1.20 1.95 1.20 0.47 1.43 3.37 2.20
 3.00 3.09 1.51 2.10 0.52 1.62 1.31 0.32 0.59 0.81
 2.81 1.87 1.18 1.35 4.75 2.48 0.96 1.89 0.90 2.05

Table 2. The behaviour of our proposed Harvda-Charvat's measure, when α is fixed at 3.5, and β varies for the given Database.

| β | $H_\alpha^\beta(P)$ | β | $H_\alpha^\beta(P)$ | β | $H_\alpha^\beta(P)$ |
|---------|---------------------|---------|---------------------|---------|---------------------|
| 0.10 | 27.64168 | 0.20 | 27.32464 | 0.30 | 27.04876 |
| 0.11 | 27.60812 | 0.21 | 27.29520 | 0.31 | 27.02344 |
| 0.12 | 27.57498 | 0.22 | 27.26617 | 0.32 | 26.99854 |
| 0.13 | 27.54224 | 0.23 | 27.23755 | 0.33 | 26.97405 |
| 0.14 | 27.50992 | 0.24 | 27.20934 | 0.34 | 26.94997 |
| 0.15 | 27.47801 | 0.25 | 27.18155 | 0.35 | 26.92631 |
| 0.16 | 27.44651 | 0.26 | 27.15417 | 0.36 | 26.90306 |
| 0.17 | 27.41543 | 0.27 | 27.12720 | 0.37 | 26.88023 |
| 0.18 | 27.38475 | 0.28 | 27.10064 | 0.38 | 26.85781 |
| 0.19 | 27.35449 | 0.29 | 27.07450 | 0.39 | 26.83581 |

Table 3. The behaviour of our proposed Harvda-Charvat's measure when $\beta=0.02$ is fixed and α varies for the given Dataset.

| α | $H_{\alpha}^{\beta}(P)$ | α | $H_{\alpha}^{\beta}(P)$ | α | $H_{\alpha}^{\beta}(P)$ |
|----------|-------------------------|----------|-------------------------|----------|-------------------------|
| 0.10 | 26.36485 | 0.20 | 26.92631 | 0.30 | 27.22810 |
| 0.11 | 26.42747 | 0.21 | 26.96598 | 0.31 | 27.24950 |
| 0.12 | 26.49334 | 0.22 | 27.00304 | 0.32 | 27.26977 |
| 0.13 | 26.55863 | 0.23 | 27.03770 | 0.33 | 27.28900 |
| 0.14 | 26.62148 | 0.24 | 27.07018 | 0.34 | 27.30727 |
| 0.15 | 26.68105 | 0.25 | 27.10064 | 0.35 | 27.32464 |
| 0.16 | 26.73701 | 0.26 | 27.12926 | 0.36 | 27.34117 |
| 0.17 | 26.78937 | 0.27 | 27.15618 | 0.37 | 27.35693 |
| 0.18 | 26.83823 | 0.28 | 27.18155 | 0.38 | 27.37196 |
| 0.19 | 26.88380 | 0.29 | 27.20549 | 0.39 | 27.38632 |

**Figure 1.** Monotonicity of our Proposed Measure.

6. CONCLUSIONS

This study presents a new generalization of Harvda-Charvat's measure of entropy, highlighting its importance in information theory and applied mathematics. The paper explores the key properties of this novel generalization. Furthermore, a new generalized mean code-word length is introduced, and the relationship between $H_{\alpha}^{\beta}(P)$ and $L_{\alpha}^{\beta}(P)$ is established through the source coding theorems presented in this paper, and also demonstrated through the Shannon-Fano Coding Algorithm. Finally, we checked its monotonicity through a precipitation dataset through which we can better anticipate the future changes to changing climatic patterns.

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